

Quick Tour of Basic Linear Algebra and Probability Theory

CS246: Mining Massive Data Sets
Winter 2011

Outline

1 Basic Linear Algebra

2 Basic Probability Theory

Matrices and Vectors

- Matrix: A rectangular array of numbers, e.g., $A \in \mathbb{R}^{m \times n}$:

$$A = \begin{pmatrix} a_{11} & a_{12} & \dots & a_{1n} \\ a_{21} & a_{22} & \dots & a_{2n} \\ \vdots & \vdots & \vdots & \vdots \\ a_{m1} & a_{m2} & \dots & a_{mn} \end{pmatrix}$$

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- Vector: A matrix consisting of only one column (default) or one row, e.g., $x \in \mathbb{R}^n$

$$x = \begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_n \end{pmatrix}$$

Matrix Multiplication

- If $A \in \mathbb{R}^{m \times n}$, $B \in \mathbb{R}^{n \times p}$, $C = AB$, then $C \in \mathbb{R}^{m \times p}$:

$$C_{ij} = \sum_{k=1}^n A_{ik} B_{kj}$$

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- Special cases: Matrix-vector product, inner product of two vectors. e.g., with $x, y \in \mathbb{R}^n$:

$$x^T y = \sum_{i=1}^n x_i y_i \in \mathbb{R}$$

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- Block multiplication: If $A = [A_{ik}]$, $B = [B_{kj}]$, where A_{ik} 's and B_{kj} 's are matrix blocks, and the number of columns in A_{ik} is equal to the number of rows in B_{kj} , then $C = AB = [C_{ij}]$ where $C_{ij} = \sum_k A_{ik} B_{kj}$

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Example: If $\vec{x} \in \mathbb{R}^n$ and $A = [\vec{a}_1 | \vec{a}_2 | \dots | \vec{a}_n] \in \mathbb{R}^{m \times n}$,

$B = [\vec{b}_1 | \vec{b}_2 | \dots | \vec{b}_p] \in \mathbb{R}^{n \times p}$:

$$A\vec{x} = \sum_{i=1}^n x_i \vec{a}_i$$

$$AB = [A\vec{b}_1 | A\vec{b}_2 | \dots | A\vec{b}_p]$$

Operators and properties

- Transpose: $A \in \mathbb{R}^{m \times n}$, then $A^T \in \mathbb{R}^{n \times m}$: $(A^T)_{ij} = A_{ji}$

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 - $tr(A) = tr(A^T)$
 - $tr(A + B) = tr(A) + tr(B)$
 - $tr(\lambda A) = \lambda tr(A)$
 - If AB is a square matrix, $tr(AB) = tr(BA)$

Special types of matrices

- Identity matrix: $I = I_n \in \mathbb{R}^{n \times n}$:

$$I_{ij} = \begin{cases} 1 & i=j, \\ 0 & \text{otherwise.} \end{cases}$$

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- Symmetric matrices: $A \in \mathbb{R}^{n \times n}$ is symmetric if $A = A^T$.
- Orthogonal matrices: $U \in \mathbb{R}^{n \times n}$ is orthogonal if $UU^T = I = U^T U$

Linear Independence and Rank

- A set of vectors $\{x_1, \dots, x_n\}$ is linearly independent if $\nexists \{\alpha_1, \dots, \alpha_n\}: \sum_{i=1}^n \alpha_i x_i = 0$

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- Properties:
 - $\text{rank}(A) \leq \min\{m, n\}$
 - $\text{rank}(A) = \text{rank}(A^T)$
 - $\text{rank}(AB) \leq \min\{\text{rank}(A), \text{rank}(B)\}$
 - $\text{rank}(A + B) \leq \text{rank}(A) + \text{rank}(B)$

Matrix Inversion

- If $A \in \mathbb{R}^{n \times n}$, $\text{rank}(A) = n$, then the inverse of A , denoted A^{-1} is the matrix that: $AA^{-1} = A^{-1}A = I$
- Properties:
 - $(A^{-1})^{-1} = A$
 - $(AB)^{-1} = B^{-1}A^{-1}$
 - $(A^{-1})^T = (A^T)^{-1}$

Range and Nullspace of a Matrix

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- $Proj(y, A) = A(A^T A)^{-1} A^T y$
- Nullspace: $null(A) = \{x \in \mathbb{R}^n \mid Ax = 0\}$

Determinant

- $A \in \mathbb{R}^{n \times n}$, $\mathbf{a}_1, \dots, \mathbf{a}_n$ the rows of A ,
 $S = \{ \sum_{i=1}^n \alpha_i \mathbf{a}_i \mid 0 \leq \alpha_i \leq 1 \}$, then $\det(A)$ is the volume of S .
- Properties:
 - $\det(I) = 1$
 - $\det(\lambda A) = \lambda \det(A)$
 - $\det(A^T) = \det(A)$
 - $\det(AB) = \det(A)\det(B)$
 - $\det(A) \neq 0$ if and only if A is invertible.
 - If A invertible, then $\det(A^{-1}) = \det(A)^{-1}$

Quadratic Forms and Positive Semidefinite Matrices

- $A \in \mathbb{R}^{n \times n}$, $x \in \mathbb{R}^n$, $x^T Ax$ is called a quadratic form:

$$x^T Ax = \sum_{1 \leq i, j \leq n} A_{ij} x_i x_j$$

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$$x^T A x = \sum_{1 \leq i, j \leq n} A_{ij} x_i x_j$$

- A is positive definite if $\forall x \in \mathbb{R}^n : x^T A x > 0$
- A is positive semidefinite if $\forall x \in \mathbb{R}^n : x^T A x \geq 0$
- A is negative definite if $\forall x \in \mathbb{R}^n : x^T A x < 0$
- A is negative semidefinite if $\forall x \in \mathbb{R}^n : x^T A x \leq 0$

Eigenvalues and Eigenvectors

- $A \in \mathbb{R}^{n \times n}$, $\lambda \in \mathbb{C}$ is an eigenvalue of A with the corresponding eigenvector $x \in \mathbb{C}^n$ ($x \neq 0$) if:

$$Ax = \lambda x$$

- eigenvalues: the n possibly complex roots of the polynomial equation $\det(A - \lambda I) = 0$, and denoted as $\lambda_1, \dots, \lambda_n$
- Properties:
 - $\text{tr}(A) = \sum_{i=1}^n \lambda_i$
 - $\det(A) = \prod_{i=1}^n \lambda_i$
 - $\text{rank}(A) = |\{1 \leq i \leq n \mid \lambda_i \neq 0\}|$

Matrix Eigendecomposition

- $A \in \mathbb{R}^{n \times n}$, $\lambda_1, \dots, \lambda_n$ the eigenvalues, and x_1, \dots, x_n the eigenvectors. $X = [x_1 | x_2 | \dots | x_n]$, $\Lambda = \text{diag}(\lambda_1, \dots, \lambda_n)$, then $AX = X\Lambda$.
- A called diagonalizable if X invertible: $A = X\Lambda X^{-1}$
- If A symmetric, then all eigenvalues real, and X orthogonal (hence denoted by $U = [u_1 | u_2 | \dots | u_n]$):

$$A = U\Lambda U^T = \sum_{i=1}^n \lambda_i u_i u_i^T$$

- A special case of Singular Value Decomposition

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2 Basic Probability Theory

Elements of Probability

- Sample Space Ω : Set of all possible outcomes
- Event Space \mathcal{F} : A family of subsets of Ω
- Probability Measure: Function $P : \mathcal{F} \rightarrow \mathbb{R}$ with properties:
 - 1 $P(A) \geq 0$ ($\forall A \in \mathcal{F}$)
 - 2 $P(\Omega) = 1$
 - 3 A_i 's disjoint, then $P(\bigcup_i A_i) = \sum_i P(A_i)$

Conditional Probability and Independence

- For events A, B :

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

- A, B independent if $P(A|B) = P(A)$ or equivalently:
 $P(A \cap B) = P(A)P(B)$

Random Variables and Distributions

- A random variable X is a function $X : \Omega \rightarrow \mathbb{R}$
Example: Number of heads in 20 tosses of a coin
- Probabilities of events associated with random variables defined based on the original probability function. e.g.,
$$P(X = k) = P(\{\omega \in \Omega \mid X(\omega) = k\})$$
- Cumulative Distribution Function (CDF) $F_X : \mathbb{R} \rightarrow [0, 1]$:
$$F_X(x) = P(X \leq x)$$
- Probability Mass Function (pmf): X discrete then
$$p_X(x) = P(X = x)$$
- Probability Density Function (pdf): $f_X(x) = dF_X(x)/dx$

Properties of Distribution Functions

■ CDF:

- $0 \leq F_X(x) \leq 1$
- F_X monotone increasing, with $\lim_{x \rightarrow -\infty} F_X(x) = 0$,
 $\lim_{x \rightarrow \infty} F_X(x) = 1$

■ pmf:

- $0 \leq p_X(x) \leq 1$
- $\sum_x p_X(x) = 1$
- $\sum_{x \in A} p_X(x) = p_X(A)$

■ pdf:

- $f_X(x) \geq 0$
- $\int_{-\infty}^{\infty} f_X(x) dx = 1$
- $\int_{x \in A} f_X(x) dx = P(X \in A)$

Expectation and Variance

- Assume random variable X has pdf $f_X(x)$, and $g : \mathbb{R} \rightarrow \mathbb{R}$.
Then

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)f_X(x)dx$$

- for discrete X , $E[g(X)] = \sum_x g(x)p_X(x)$

- Properties:

- for any constant $a \in \mathbb{R}$, $E[a] = a$

- $E[ag(X)] = aE[g(X)]$

- Linearity of Expectation:

$$E[g(X) + h(X)] = E[g(X)] + E[h(X)]$$

- $Var[X] = E[(X - E[X])^2]$

Some Common Random Variables

- $X \sim \text{Bernoulli}(p)$ ($0 \leq p \leq 1$):

$$p_X(x) = \begin{cases} p & x=1, \\ 1-p & x=0. \end{cases}$$

- $X \sim \text{Geometric}(p)$ ($0 \leq p \leq 1$): $p_X(x) = p(1-p)^{x-1}$

- $X \sim \text{Uniform}(a, b)$ ($a < b$):

$$f_X(x) = \begin{cases} \frac{1}{b-a} & a \leq x \leq b, \\ 0 & \text{otherwise.} \end{cases}$$

- $X \sim \text{Normal}(\mu, \sigma^2)$:

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2\sigma^2}(x-\mu)^2}$$

Multiple Random Variables and Joint Distributions

X_1, \dots, X_n random variables

■ Joint CDF: $F_{X_1, \dots, X_n}(x_1, \dots, x_n) = P(X_1 \leq x_1, \dots, X_n \leq x_n)$

■ Joint pdf: $f_{X_1, \dots, X_n}(x_1, \dots, x_n) = \frac{\partial^n F_{X_1, \dots, X_n}(x_1, \dots, x_n)}{\partial x_1 \dots \partial x_n}$

■ Marginalization:

$$f_{X_1}(x_1) = \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f_{X_1, \dots, X_n}(x_1, \dots, x_n) dx_2 \dots dx_n$$

■ Conditioning: $f_{X_1|X_2, \dots, X_n}(x_1|x_2, \dots, x_n) = \frac{f_{X_1, \dots, X_n}(x_1, \dots, x_n)}{f_{X_2, \dots, X_n}(x_2, \dots, x_n)}$

■ Chain Rule: $f(x_1, \dots, x_n) = f(x_1) \prod_{i=2}^n f(x_i|x_1, \dots, x_{i-1})$

■ Independence: $f(x_1, \dots, x_n) = \prod_{i=1}^n f(x_i)$.

■ More generally, events A_1, \dots, A_n independent if

$$P(\bigcap_{i \in S} A_i) = \prod_{i \in S} P(A_i) \quad (\forall S \subseteq \{1, \dots, n\}).$$

Random Vectors

X_1, \dots, X_n random variables. $X = [X_1 X_2 \dots X_n]^T$ random vector.

- If $g : \mathbb{R}^n \rightarrow \mathbb{R}$, then

$$E[g(X)] = \int_{\mathbb{R}^n} g(x_1, \dots, x_n) f_{X_1, \dots, X_n}(x_1, \dots, x_n) dx_1 \dots dx_n$$

- if $g : \mathbb{R}^n \rightarrow \mathbb{R}^m$, $g = [g_1 \dots g_m]^T$, then

$$E[g(X)] = [E[g_1(X)] \dots E[g_m(X)]]^T$$

- Covariance Matrix:

$$\Sigma = \text{Cov}(X) = E[(X - E[X])(X - E[X])^T]$$

- Properties of Covariance Matrix:

- $\Sigma_{ij} = \text{Cov}[X_i, X_j] = E[(X_i - E[X_i])(X_j - E[X_j])]$
- Σ symmetric, positive semidefinite

Multivariate Gaussian Distribution

$\mu \in \mathbb{R}^n$, $\Sigma \in \mathbb{R}^{n \times n}$ symmetric, positive semidefinite
 $X \sim \mathcal{N}(\mu, \Sigma)$ n -dimensional Gaussian distribution:

$$f_X(x) = \frac{1}{(2\pi)^{n/2} \det(\Sigma)^{1/2}} \exp\left(-\frac{1}{2}(x - \mu)^T \Sigma^{-1}(x - \mu)\right)$$

- $E[X] = \mu$
- $Cov(X) = \Sigma$

Parameter Estimation: Maximum Likelihood

Parametrized distribution $f_X(x; \theta)$ with parameter(s) θ unknown.

i.i.d. samples x_1, \dots, x_n observed.

Goal: Estimate θ

MLE: $\hat{\theta} = \operatorname{argmax}_{\theta} \{f(x_1, \dots, x_n; \theta)\}$

MLE Example

$X \sim \text{Gaussian}(\mu, \sigma^2)$. $\theta = (\mu, \sigma^2)$ unknown. Samples x_1, \dots, x_n .
Then:

$$f(x_1, \dots, x_n; \mu, \sigma^2) = \left(\frac{1}{2\pi\sigma^2}\right)^{n/2} \exp\left(-\frac{\sum_{i=1}^n (x_i - \mu)^2}{2\sigma^2}\right)$$

Setting: $\frac{\partial \log f}{\partial \mu} = 0$ and $\frac{\partial \log f}{\partial \sigma} = 0$

Gives:

$$\hat{\mu}_{MLE} = \frac{\sum_{i=1}^n x_i}{n}, \quad \hat{\sigma}_{MLE}^2 = \frac{\sum_{i=1}^n (x_i - \hat{\mu})^2}{n}$$

If not possible to find the optimal point in closed form, iterative methods such as gradient decent can be used.

Some Useful Inequalities

- Markov's Inequality: X random variable, and $a > 0$. Then:

$$P(|X| \geq a) \leq \frac{E[|X|]}{a}$$

- Chebyshev's Inequality: If $E[X] = \mu$, $\text{Var}(X) = \sigma^2$, $k > 0$, then:

$$\Pr(|X - \mu| \geq k\sigma) \leq \frac{1}{k^2}$$

- Chernoff bound: X_1, \dots, X_n iid random variables, with $E[X_i] = \mu$, $X_i \in \{0, 1\}$ ($\forall 1 \leq i \leq n$). Then:

$$P\left(\left|\frac{1}{n} \sum_{i=1}^n X_i - \mu\right| \geq \epsilon\right) \leq 2 \exp(-2n\epsilon^2)$$

- Multiple variants of Chernoff-type bounds exist, which can be useful in different settings

References

- 1 CS229 notes on basic linear algebra and probability theory
- 2 Wikipedia!